

Basic Radar 3.1: Probability Theory for Incoherent Scatter Radar

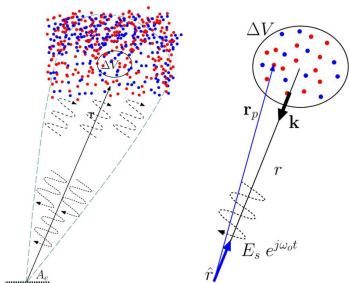
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July, 2020

The Need for Statistical Descriptions of ISR Signals

If I knew the positions of every single electron in the scattering volume, I would know the received voltage exactly:



Exact expression for scattered electric field as a superposition of Thompson scatterers:

$$E_s = -\frac{r_e}{r} E_0 \sum_{p=1}^{N_0 \Delta V} e^{jk \cdot r_p}$$

ISR theory predicts statistical aspects of the scattered signal:

Scattered Power: $\langle |E_s|^2 \rangle$ Autocorrelation Function: $\langle E_s(t) E_s^*(t - \tau) \rangle$

These statistical properties are functions of macroscopic properties of the plasma: N_e , T_e , T_i , u_{los} .

Random Variables

A **random variable** is a variable whose numerical value depends on the outcome of a probabilistic phenomenon.

Probability Density Function:

$$P(x_1 < X < x_2) = \int_{x_1}^{x_2} p_X(x) dx$$

Expected Values:

$$E\{g(X)\} = \int_{-\infty}^{\infty} g(x)p_X(x) dx$$

Mean:

$$\text{Mean}\{X\} = E\{X\} = \bar{X}$$

Variance:

$$\text{Var}\{X\} = E\{(X - E\{X\})^2\} = E\{X^2\} - (E\{X\})^2$$

Collections of Random Variables

Multiple RVs must be described by joint-PDFs:

$$P(x_0 < X < x_1 \cup y_0 < Y < y_1) = \int_{x_0}^{x_1} \int_{y_0}^{y_1} p_{XY}(x, y) dy dx$$

If X and Y are **independent**:

$$p_{XY}(x, y) = p_X(x)p_Y(y) \quad p_{X|Y}(x|y) = p_X(x)$$

Relationships between RVs are defined through covariances:

$$\text{Cov}\{X, Y\} = E\{(X - E\{X\})(Y - E\{Y\})\}$$

Uncorrelated RVs have $\text{Cov}\{X, Y\} = 0$

Independent RVs are uncorrelated, but uncorrelated RVs are not necessarily independent.

Random Vectors

Column vector of random variables

$$\mathbf{X} = \begin{pmatrix} X_0 \\ X_1 \\ \vdots \\ X_{N-1} \end{pmatrix}$$

Covariance matrix of a random vector

$$\begin{aligned} K_{\mathbf{X}} &= \text{Cov} \{ \mathbf{X} \} = E \{ (\mathbf{X} - \bar{\mathbf{X}}) (\mathbf{X} - \bar{\mathbf{X}})^T \} \\ &= \begin{pmatrix} \text{Var} \{ X_0 \} & \text{Cov} \{ X_0, X_1 \} & \cdots & \text{Cov} \{ X_0, X_{N-1} \} \\ \text{Cov} \{ X_1, X_0 \} & \text{Var} \{ X_1 \} & \cdots & \text{Cov} \{ X_1, X_{N-1} \} \\ \vdots & \vdots & \ddots & \vdots \\ \text{Cov} \{ X_{N-1}, X_0 \} & \text{Cov} \{ X_{N-1}, X_1 \} & \cdots & \text{Var} \{ X_{N-1} \} \end{pmatrix} \end{aligned}$$

Cross-covariance of two random vectors

$$K_{\mathbf{X}\mathbf{Y}} = \text{Cov} \{ \mathbf{X}, \mathbf{Y} \} = E \{ (\mathbf{X} - \bar{\mathbf{X}}) (\mathbf{Y} - \bar{\mathbf{Y}})^T \}$$

Gaussian Distribution

A Gaussian random variable X has the following probability density function (Normal Distribution):

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{x - \mu}{2\sigma^2}\right\}$$
$$E\{X\} = \mu \quad \text{Var}\{X\} = \sigma^2$$
$$E\{(X - \mu)^4\} = 3\sigma^4$$

A jointly-Gaussian vector of random variables $\mathbf{X} = [X_0, X_1, X_2, \dots, X_{N-1}]^T$ has the joint pdf:

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{N}{2}} |K|^{\frac{1}{2}}} \exp\left\{-\frac{1}{2} [\mathbf{x} - \boldsymbol{\mu}]^T K^{-1} [\mathbf{x} - \boldsymbol{\mu}]\right\}$$
$$E\{\mathbf{X}\} = \boldsymbol{\mu}$$
$$\text{Cov}\{\mathbf{X}\} = E\{[\mathbf{X} - \boldsymbol{\mu}][\mathbf{X} - \boldsymbol{\mu}]^T\} = K$$

Central Limit Theorem

Given a set of finite-variance, independent and identically distributed RV, $[X_0, X_1, \dots, X_{K-1}]$, the distribution function of the average:

$$\hat{X} = \frac{1}{K} \sum_{n=0}^{K-1} X_n$$

will asymptotically approach a Gaussian distribution as K increases.

$$E \{ \hat{X} \} = E \{ X_n \} \quad \text{Var} \{ \hat{X} \} = \frac{1}{K} \text{Var} \{ X_n \}$$

This is an amazingly useful theorem:

- Only the mean and variances of the intermediate quantities need to be calculated to predict the distribution of the final averaged result.
- Distribution functions of intermediate quantities do not need to be calculated in detail since the final averaged result will just be Gaussian.

Properties of Jointly Gaussian Random Variables

- Linear combinations:

$$Z = \alpha X + \beta Y + \gamma \quad E\{Z\} = \alpha E\{X\} + \beta E\{Y\} + \gamma$$
$$\text{Var}\{Z\} = \alpha^2 \text{Var}\{X\} + \beta^2 \text{Var}\{Y\} + 2\alpha\beta \text{Cov}\{X, Y\}$$

- Matrix generalization:

$$Y = AX + b \quad E\{Y\} = AE\{X\} + b \quad \text{Cov}\{Y\} = A\text{Cov}\{X\}A^T$$

Special cases for zero mean random variables:

- Odd moments are zero:

$$E\{V_1\} = E\{V_1 V_2 V_3\} = E\{V_1 V_2 V_3 V_4 V_5\} = \dots = 0$$

- Fourth moment theorem: $E\{V_1 V_2 V_3 V_4\} =$

$$E\{V_1 V_2\} E\{V_3 V_4\} + E\{V_1 V_3\} E\{V_2 V_4\} + E\{V_1 V_4\} E\{V_2 V_3\}$$

- General even moment theorem (Isserlis' Theorem)

$$E\{V_1 V_2 \dots V_{2n-1} V_{2n}\} = \sum \prod E\{V_i V_j\}$$

Complex-valued Random Variables

A complex-valued random variable X can be described by

$$\text{Mean: } \bar{X} = E \{X\}$$

$$\text{Covariance: } K_X = E \{ (X - \bar{X}) (X - \bar{X})^* \}$$

$$\text{Pseudo-Covariance: } J_X = E \{ (X - \bar{X}) (X - \bar{X}) \}$$

A vector of complex-valued random variables \mathbf{X} can be described by

$$\text{Mean: } \bar{\mathbf{X}} = E \{\mathbf{X}\}$$

$$\text{Covariance: } K_{\mathbf{X}} = E \{ (\mathbf{X} - \bar{\mathbf{X}}) (\mathbf{X} - \bar{\mathbf{X}})^H \}$$

$$\text{Pseudo-Covariance: } J_{\mathbf{X}} = E \{ (\mathbf{X} - \bar{\mathbf{X}}) (\mathbf{X} - \bar{\mathbf{X}})^T \}$$

Where $(\cdot)^H$ means Hermitian transpose (i.e. with a complex conjugate) and $(\cdot)^T$ means ordinary transpose.

Relationship to Real Random Variables

A vector of N complex-valued random variables can be written as two vectors of real-valued random variables representing the real and imaginary parts.

$$\mathbf{X} = \mathbf{X}_R + j\mathbf{X}_I$$

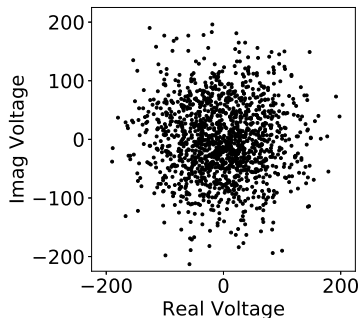
The covariance and cross-covariance matrices of these real vectors are related to the covariance and pseudo-covariance of the complex vector by

$$\begin{aligned}\text{Cov} \{ \mathbf{X}_R \} &= \frac{1}{2} \Re K_X + \frac{1}{2} \Re J_X \\ \text{Cov} \{ \mathbf{X}_I, \mathbf{X}_R \} &= \frac{1}{2} \Im K_X + \frac{1}{2} \Im J_X \\ \text{Cov} \{ \mathbf{X}_R, \mathbf{X}_I \} &= -\frac{1}{2} \Im K_X + \frac{1}{2} \Im J_X \\ \text{Cov} \{ \mathbf{X}_I \} &= \frac{1}{2} \Re K_X - \frac{1}{2} \Re J_X\end{aligned}$$

Special Properties of ISR Voltages

ISR Voltages will always be:

- Gaussian
- Zero Mean: ($E\{V\} = 0$)
- Finite power: ($E\{VV^*\} < \infty$)
- Random Phase:
 - Zero Pseudo-variance ($E\{VV\} = 0$)
 - $\text{Cov}\{\Re V, \Im V\} = 0$
- Collections of ISR voltages will always have zero pseudo-covariance with each other.



With these properties, the complex-valued covariances between ISR voltages tells us everything we could want to know.

Describing ISR Voltages

ISR signals are complex valued, zero mean, and random phase.

$$V = V_R + jV_I \quad E\{V_R\} = E\{V_I\} = 0$$

$$E\{VV^*\} = \sigma^2 \quad E\{V_R V_I\} = 0 \quad \text{Cov} \left\{ \begin{pmatrix} V_R \\ V_I \end{pmatrix} \right\} = \frac{1}{2} \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix}$$

When we talk about correlations between ISR signals

$$E\{V_1 V_1^*\} = \sigma_1^2 \quad E\{V_2 V_2^*\} = \sigma_2^2$$

$$E\{V_1 V_2^*\} = \rho = \rho_R + j\rho_I$$

What we really mean is

$$V_1 = V_{1R} + jV_{1I} \quad V_2 = V_{2R} + jV_{2I}$$
$$\text{Cov} \left\{ \begin{pmatrix} V_{1R} \\ V_{1I} \\ V_{2R} \\ V_{2I} \end{pmatrix} \right\} = \frac{1}{2} \begin{pmatrix} \sigma_1^2 & 0 & \rho_R & -\rho_I \\ 0 & \sigma_1^2 & \rho_I & \rho_R \\ \rho_R & \rho_I & \sigma_2^2 & 0 \\ -\rho_I & \rho_R & 0 & \sigma_2^2 \end{pmatrix}$$

Probability for ISR Summary

- The theory of ISR is a probabilistic theory for the statistical properties of the received voltages.
- ISR voltages are Gaussian, zero mean, random phase, complex-valued random variables.
- All the information we want is in the variances (power) and covariances between voltages.